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() (χ^2)

Optimum Constant Smoothing for Exponential Smoothing Model with Application

ABSTRACT

In this paper, we study methods to obtain the constant smoothing of exponential smoothing model and suggest another method, to obtain constant smoothing of dependent (two estimation New and old Information) of dependent adaptive filtering, and applied for time series to National Income overall of the Egypt for (1965-2002) and unsure of use this model to use the measure of (χ^2), simulation have been done depending on the parameter of the original series with different numbers.

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2007/5/27 :

2007/3/1 :

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(())

(Difference)

(...)

Baraned)

(1959

Adaptive)

(Filtering

() -3 Stanley.L.S. -2 -1

_____:

(Holt.C.C) (1958)

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$$S_t = \alpha Y_t + (1 - \alpha)S_{t-1} \tag{1}$$

(S_{t-1}) (Smoothed Statistic) $:(S_t)$
 (α) $(t - 1)$
 (x_t) (Smoothing Constant)
 (S_t) (S_{t-1})

$$S_t = \alpha y_t + (1 - \alpha)[\alpha y_{t-1} + (1 - \alpha)S_{t-2}]$$

$$S_t = \alpha y_t + \alpha(1 - \alpha)y_{t-1} + (1 - \alpha)^2 S_{t-2} \tag{2}$$

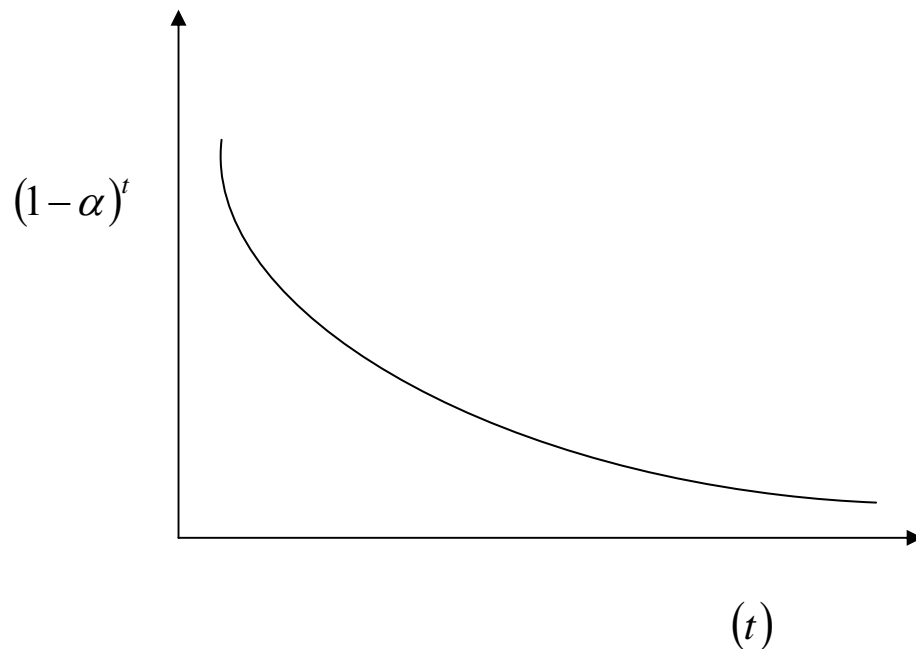
$$S_t = \alpha Y_t + \alpha(1 - \alpha)y_{t-1} + \alpha(1 - \alpha)^2 y_{t-2} + \alpha(1 - \alpha)^3 y_{t-3} + \alpha(1 - \alpha)^4 y_{t-4} + \dots + (1 - \alpha)^t S_0 \tag{3}$$

$$S'_t = \alpha \sum_{r=0}^{t-1} (1 - \alpha)^r y_{t-r} + (1 - \alpha)^t y_0 \tag{4}$$

$$:(1 - \alpha)^r \quad ; So \quad ;$$

$$(0.1 < \alpha < 0.5)$$

(James,2003)



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Specifying the Smoothing Weights

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(Smoothing Constant)

و $(0.1 < \alpha < 0.5)$

$$F_{t+1} = \alpha y_t + (1 - \alpha)F_t \quad (5)$$

Simple Exponential)

(α)

(Smoothing

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C.C.Holt 1958

Browns ,1963

Harrison 1965

.(Spyros.M &els, 1983) :

X_{t-N}

X_{t-N}

$$F_{t+1} = F_t + \left(\frac{x_t}{N} - \frac{x_{t-N}}{N} \right) \quad (6)$$

F_t

$$F_{t+1} = F_t + \left(\frac{x_t}{N} - \frac{F_t}{N} \right) \quad (7)$$

:

$$F_{t+1} = \left(\frac{1}{N} \right) x_t + \left(1 - \frac{1}{N} \right) F_t \quad (8)$$

$$(1/N) \quad (F_{t+1}) \quad (3)$$

$$(1-1/N) \quad F_t$$

$$N \quad) \quad (1/N) \quad (N)$$

$$(1/N) \quad (\alpha) \quad (N=1) \quad) \quad ($$

: (8)

$$F_{t+1} = \alpha y_t + (1 - \alpha) F_t \quad (9)$$

Kendall

Stanley _____ **-2**

(1976)

(2002)

$$Y_t = \beta_t + e_t \tag{10}$$

$$\beta_t \quad \sigma_e^2 \quad e_t$$

$$\beta_t = \beta_{t-1} + \delta_t \tag{11}$$

$$(e_t) \quad (\delta_t)$$

$$(\delta_t) \quad (\beta_t)$$

$$: \quad (Y_t) \quad (\beta_0)$$

$$Y_t = \beta_t + e_t \tag{12}$$

$$Y_{t-1} = \beta_{t-1} + e_{t-1} \tag{13}$$

$$(11) \quad (13) \quad (12)$$

(15)

$$Y_t - Y_{t-1} = \beta_t - \beta_{t-1} + e_t - e_{t-1} \tag{14}$$

$$= \delta_t + e_t - e_{t-1}$$

$$= (\delta_t + e_t) - e_{t-1} \tag{15}$$

$$= u_t + \theta u_{t-1}, \tag{16}$$

$$u_t = \delta_t + e_t \tag{17}_t$$

$$e_{t-1} \quad \theta u_{t-1} \quad \theta$$

(17)

$$\sigma_u^2 = \sigma_\delta^2 + \sigma_e^2 \tag{18}$$

$e_t \quad \delta_t$,

$$\sigma_e^2 = \text{var}(e_{t-1}) = \text{var}(\theta u_{t-1}) = \theta^2 \text{var}(u_{t-1}) = \theta^2 \sigma_u^2 = \theta^2 (\sigma_\delta^2 + \sigma_e^2) \tag{19}$$

(20)

$$\sigma_e^2 = \theta^2(\sigma_\delta^2 + \sigma_e^2) \quad (20)$$

(21)

(20)

$$\theta^2 = \frac{\sigma_e^2}{\sigma_\delta^2 + \sigma_e^2} \quad (21)$$

 θ^2 β_t σ_δ^2

.

$$Y_t - Y_{t-1} = u_t - \theta u_{t-1} \quad (22)$$

$$Y_t = Y_{t-1} + u_t - \theta u_{t-1} \quad (23)_t$$

 Y_t Y_{t-1}

$$E(Y_t / Y_{t-1}) = E[Y_{t-1} + u_t - \theta u_{t-1} / Y_{t-1}] \quad (24)$$

$$= Y_{t-1} + E[u_t / Y_{t-1}] - E[\theta u_{t-1} / Y_{t-1}] \quad (25)$$

$$E(Y_t / Y_{t-1}) = Y_{t-1} - \theta u_{t-1} \quad (26)$$

(23)

(26)

$$Y_t - E(Y_t / Y_{t-1}) = (Y_{t-1} + u_t - \theta u_{t-1}) - (Y_{t-1} - \theta u_{t-1}) = u_t \quad (27)$$

t

(27)

(26)

$$E(Y_t/Y_{t-1}) = Y_{t-1} - \theta u_{t-1} \tag{28}$$

$$(28) \qquad (27)$$

$$E(Y_t/Y_{t-1}) = Y_{t-1} - \theta [Y_{t-1} - E(Y_{t-1}/Y_{t-2})] \tag{29}$$

θ

Y_{t-1}

$$E(Y_t/Y_{t-1}) = (1-\theta)Y_{t-1} + \theta E(Y_{t-1}/Y_{t-2})_t \tag{30}$$

(t-1)

t

$E[Y_{t-1}/Y_{t-2}]$

Y_{t-1}

(1- θ)

(θ)

. (1- θ)

:(_____) _____ -

1971

Stevens

Harrisons

(Kalman filter)

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$$\begin{array}{ccc} Y_t & t & F_t \\ & t+1 & \end{array}$$

$$F_{t+1} = \alpha Y_t + (1 - \alpha) F_t \quad (31)$$

وان F_t و Y_t على فرض انهما مستقلان ، واذا كان تباين Y_t هو σ_y^2 وتباين F_t هو σ_F^2 وبهذه الحالة ممكن ان نحسب التباين الكلي للدالة او التباين الكلي الموزون كما هو ادناه

$$\sigma_{F_{t+1}}^2 = (1 - \alpha)^2 \sigma_F^2 + \alpha^2 \sigma_y^2 \quad (32)$$

$$(31) \quad \alpha$$

$$(32)$$

$$\frac{\partial \sigma_{F_{t+1}}^2}{\partial \alpha} = -2(1 - \alpha) \sigma_F^2 + 2\alpha \sigma_y^2 = 0$$

α

$$2(1 - \alpha) \sigma_F^2 = 2\alpha \sigma_y^2$$

$$\sigma_F^2 - \alpha \sigma_F^2 = \alpha \sigma_y^2$$

$$\Rightarrow \sigma_F^2 = \alpha \sigma_y^2 + \alpha \sigma_F^2$$

$$\alpha = \frac{\sigma_F^2}{\sigma_F^2 + \sigma_y^2} \quad (33)$$

$$(33)$$

$$(31)$$

α

$$F_{t+1} = \frac{\sigma_F^2}{\sigma_F^2 + \sigma_y^2} Y_t + \left(1 - \frac{\sigma_F^2}{\sigma_F^2 + \sigma_y^2}\right) F_t \quad (33)$$

Y_t

F_t

$$\sigma_{F_{t+1}}^2 = \left(1 - \frac{\sigma_F^2}{\sigma_F^2 + \sigma_y^2}\right)^2 \sigma_F^2 + \left(\frac{\sigma_F^2}{\sigma_F^2 + \sigma_y^2}\right)^2 \sigma_y^2 \quad (34)$$

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(70) -1

(10) (10))

) ((10)

((

(100) -2

(20) (20) (20))

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	Stanley.LS	()	
:			-1
		(1/N)	
	N	(1-1/N)	F _t
(N)	(1/N)
	(1/N)	α	(N=1)
		Stanley.LS	-2
(β)		(β)	
			-3
			-4
			-5

(2006)

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